# Precipitous Declines in Northwest Atlantic Dusky Sharks 

Author list

## Supplementary Information

## Data Sources

We attempted to consider all long-term series which satisfied the following conditions. First, we considered surveys in which methods had remained standard. For example, in many surveys for sharks there has been a change from wire to monofilement leaders, which makes it more difficult to distinguish long-term trends. Second, we only used data in which sharks were examined by scientifically trained observers. This included the one commercial series, the Crooke data series, in which the fishermen, Mr. Crooke, were trained by scientists at the Mote Marine Laboratory (Hueter 1991). Third, we only used data in which sharks were brought to very close contact with the scientifically trained observers, usually on board the vessel. This last step was considered necessary because of the difficulty of distinguishing dusky sharks from similar species, such as the sandbar shark.

Specifically, data on dusky shark abundance was obtained from four fishery-independent bottom trawl surveys (National Marine Fishery Service (NMFS) northeast U.S. offshore bottom trawl survey: 1966-2005, NMFS northeast U.S. inshore bottom trawl survey: 1974-2004, southeast U.S. SEAMAP bottom shrimp trawl survey: 1989-2003, northern

longline survey [North Carolina Institute of Marine Sciences longline survey (NC IMS): 1972-2002] and one commercial longline observer data set [Crooke: 1976-1989 (Hueter 1991)]. These data were collected along the coastal waters of east coast of the U.S. and the northern Gulf of Mexico, from the Gulf of Maine to Louisiana (Fig X - map showing area of all surveys). Although the NMFS offshore data set extends back to 1963, we limited the analysis of the data set to 1966 onward which maintained consistent spatial coverage. As well, after 199X, deeper water SEAMAP strata were no longer sampled. We did not consider these strata in our analysis.

## Estimating trends in abundance

## General modeling strategy

Trends in relative abundance of dusky sharks from each survey were analyzed using generalized linear models with a negative binomial error structure and a log link. The negative binomial error structure is appropriate for data with a large number of zero (no catch) observations. The log link allows the long-term trend in relative abundance to be characterized. All analyses were conducted using SAS v9.1 (SAS Institute Inc., Cary, NC, USA). The probability of catching $C_{i}$ individuals of a given species in survey tow $i$ was assumed to follow a negative binomial distribution with the mean $\mu_{i}$,
$p\left(C_{i} ; k ; \mu_{i}\right)=\frac{\Gamma\left(C_{i}+\frac{1}{k}\right)}{\Gamma\left(C_{i}+1\right) \Gamma\left(\frac{1}{k}\right)} \frac{\left(k \mu_{i}\right)^{k}}{\left(1+k \mu_{i}\right)^{C_{i}+\left(\frac{1}{k}\right)}}$, for $C_{i}=0,1,2, \ldots$,
where $\Gamma$ is the gamma function and $k$ is the negative binomial dispersion parameter. The expected mean catch of a given species is then,
$\log \left(\mu_{i}\right)=\boldsymbol{x}_{\boldsymbol{i}}^{\prime} \boldsymbol{\beta}+\log (\delta)$
where $\boldsymbol{x}_{\boldsymbol{i}}^{\prime}$ is a vector of explanatory covariates for observation $i, \boldsymbol{\beta}$ is a vector of unknown coefficients for the explanatory variables and $\delta$ is the offset term.

## Trawl surveys

For trawl surveys, considerable ancillary data existed which could be used as meaningful explanatory covariates. In general, these included temperature, depth, geographic location, and julian day. The NMFS surveys and the SEAMAP southeast US survey cover a large range of latitudes that are sampled up to three times each year. At the same time, dusky shark is known to undertake large north-south migrations throughout the year (Bonfil, R. Fish. Res. 29, 101-117 (1997)). To account for this, we assumed that the expected catch was dependent on the day of year (seasonal cycle), latitude and the interaction of the two. The seasonal cycle, $q$, was characterized by a series of sine and cosine terms, with periods, $j$, of $1 / 2$ and 1 year as,
$q\left(d_{i}\right)=\sum_{j=1}^{2}\left[\varsigma_{j} \cos \left(\frac{2 \pi j d_{i}}{365.25}\right)+\sigma_{j} \sin \left(\frac{2 \pi j d_{i}}{365.25}\right)\right]$
where $d_{i}$ is the sequential day of the year that observation $i$ occurred in, and $\varsigma_{i}$ and $\sigma_{i}$ are estimated parameters. This was included in the vector of explanatory covariates ( $\boldsymbol{x}_{\boldsymbol{i}}^{\prime}$ ) for the NMFS surveys and southeast U.S. SEAMAP survey as,
$\boldsymbol{x}_{\boldsymbol{i}}^{\prime}=y_{i}+o_{i}+o_{i}^{2}+t_{i}+t_{i}^{2}+l_{i}+q\left(d_{i}\right)+l \cdot q\left(d_{i}\right)$
where $y_{i}$ is the year that observation $i$ occurred in, $o_{i}$ is the depth, $t_{i}$ is bottom temperature and $l_{i}$ is latitude. We used this vector of covariates in equation 2 along with trawl width (NFMS surveys) or swept area (southeast U.S. SEAMAP survey) as the offset term ( $\delta$ ) to estimate the vector of parameters,

$$
\begin{equation*}
\boldsymbol{\beta}=\beta_{y}+\beta_{o}+\beta_{o^{2}}+\beta_{t}+\beta_{t^{2}}+\beta_{l}+\beta_{q(d)}+\beta_{l \cdot q(d)} \tag{5}
\end{equation*}
$$

Estimates from the northern Gulf of Mexico bottom shrimp trawl survey were previously derived (Shepherd and Myers 2005) from a model with non-significant factors removed. The final northern Gulf of Mexico model included year and depth as parameters.

## North Carolina Institute of Marine Sciences longline survey

For the NC IMS survey, ancillary data was limited to the year and seasonal cycle ( $\beta_{q(d)}$ ) and a fixed sampling station effect ( $\beta_{f}$ ), where $f$ is the fixed station identifier where observation i occurred.

## Crooke data series

The Crooke data set is from one commercial shark fisherman who was trained to identify sharks. Although this fisherman kept excellent records, there are two limitations with the data set. First, during the first years of the survey, Mr. Crooke appeared to be learning the best places and time to fish, and his catch rates may be have reflected abundance. Second, he did not record fishing sets where he did not catch any sharks. The first problem was investigated by using different start dates for the analysis (Table XX). Using different start dates did have a small effect on estimates of trends in abundance. Using a later start date resulted in larger rates of decline. However, the differences are small and any of the models are consistent with our results. ${ }^{* * *}$ We went with 1976? Why? ${ }^{* * *}$.

The second problem, i.e. no zero sets, is not as large a problem as it might first appear because dusky sharks were not a large part of the overall catch (give numbers). Thus, most of the zero catches for dusky shark are present in any case. Furthermore, as the overall catch rate declined, Mr. Crooke, increased the number of hooks fished from 15 to 20. The difficulty of the lack of records when a longline was set, and no sharks were caught (or reported) was dealt with in 3 different ways. First, a truncated negative binomial model was fit to the data, and the zeros were inputted, using the fit to the data. Then a regular negative binomial model was fit to the data. The difficulty with this approach is that it is difficult to include the uncertainty about the number of zeros in the model. Second, we fit a trend in the total number of sharks present using a truncated negative binomial model, and then fit a logistic to the proportion of dusky sharks present. This approach has the advantage that it is relatively easy to estimate the standard errors of
the overall slope estimates, because the two estimates are independent because the parameterisation is orthogonal (Welch et al. 1996). Third, for robustness, we examined the consequence of assuming that there were no missing zeros. [ *** Ram, did you do all this?]. *** Refer to table for results of this analysis *** *** say what final model was

## Model Robustness

We assessed model robustness in two ways. For data sources that included significant ancillary data (trawl surveys), we constructed a series of ecologically sensible models from various combinations of covariates. From each of these models, we calculated Akaike Information Criterion (AIC), which is a measure of model parsimony. In all but one case, the models we used (equation $4 \& 5$ ) had the lowest AIC (Table XXX). In the other case, our model AIC was only slightly higher than an alternate model. Regardless of the models used, estimates of instantaneous rate of change in abundance were similar.

We also verified the robustness of our models to assumptions of error structure and model form. [We fit delta log-normal, delta log-gamma, and zero-truncated negative binomial models(?). Results were similar for all models (Table XXX).]

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Comment [TDS5]: To what scope do I include
these results?
Comment [TDS6]: Do we do this for all surveys
or just Crooke?
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## Likelihood profiles

The generalized linear models allow estimation of the instantaneous rate of change in abundance ( $\beta_{y}$ ) and symmetrical errors. We take the approach of examining the log-
likelihood profiles of $\beta_{y}$, which allows for non-symmetrical errors and estimation in cases with few data points. [more about log-likelihoods] $\qquad$

## Effectiveness of management measures

We examined trends in abundance for change since the implementation of the NMFS Atlantic Sharks Management Plan in 1993 through the use of piece-wise generalized linear models. These models are constructed by included the additional parameters $y_{i}^{*}$ in the vector $\boldsymbol{x}_{\boldsymbol{i}}^{\prime}$ (eq. 4) and $\beta_{y}^{*}$ in the vector $\boldsymbol{\beta}$ (eq. 5), where $y_{i}{ }^{*}=0$ if $y_{i}<$ 1993 and $y_{i}{ }^{*}=y_{i}-1993$ otherwise, and $\beta_{y}{ }^{*}$ is the difference in the instantaneous rate of change in abundance before and after 1993.

Trawl vs. longline analysis
? $\qquad$ Comment [TDS8]: Was anything done here?

Change in length/weight
Lengths of captured dusky sharks were available from the NMFS trawl surveys (offshore mean $=115.4, \mathrm{se}=5.1, \min =19, \max =186)($ inshore mean $=96.8, \mathrm{se}=2.3, \min =37$, max=211), the NC IMS longline survey (mean=99.6, $\mathrm{se}=0.88, \min =30.5$, $\max =290$ ) and

## Comment [TDS9]: Too small

 the Crooke commercial data (mean=240.4, $\mathrm{se}=4.1$, $\min =185.4$, $\max =302.3$ ). We did not further consider dusky length from the trawl surveys. They appear to primarily to catch juveniles, and there is no a priori reason we expect a change in length of juvenile (*** what are the means ${ }^{* * *}$ ). Changes in length from each of the longline surveys wereestimated using generalized linear models with log links and gamma error distributions.
Estimates were: NMFS offshore (0.0110, 95\% CI:-0.0043 to $0.0262, \mathrm{SE}=0.0078$ ),

NMFS inshore (0.0057, 95\% CI:-0.0045 to 0.0159, SE= 0.0052), NC IMS (-0.0105, 95\%
CI: -0.0132 to $-0.0077, \mathrm{SE}=0.0014)$. Meta-analytic mean (-.0088, $\mathrm{SE}=.0037$ )

## References

## Table XX

| negative binomial |  | delta-gamma |  | delta-lognormal |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Year | SE | Year | SE | Year | SE |
|  |  |  |  |  |  |  |
| NMFS offshore | -0.076 | 0.035 | -0.087 | 0.045 | -0.070 | 0.051 |
| NMFS inshore | -0.105 | 0.033 | -0.091 | 0.032 | -0.147 | 0.061 |
| SEAMAP | -0.205 | 0.077 | -0.256 | 0.064 | -0.257 | 0.064 |
| NC IMS | -0.162 | 0.017 | -0.219 | 0.021 | -0.231 | 0.024 |
| Crooke | -0.191 | 0.062 | -0.250 | 0.069 | -0.267 | 0.073 |

Table XX.

| Start year | Zeros included | Estimate | StdErr | $p$ |
| :--- | :--- | :--- | :--- | :--- |
| 1975 | Y | -0.1596 | 0.0532 | 0.0027 |
| 1975 | N | -0.1493 | 0.0502 | 0.0029 |
| 1976 | Y | -0.1905 | 0.0619 | 0.0021 |
| 1976 | N | -0.1521 | 0.0666 | 0.0224 |
| 1977 | Y | -0.2350 | 0.0664 | 0.0004 |
| 1977 | N | -0.1886 | 0.0716 | 0.0085 |

Table XX.

| Survey | Model $^{1}$ | Parameters | AIC | Year | Year SE | K | K SE |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| NMFS offshore | A | 16 | 266.2 | -0.076 | 0.035 | 3.515 | 1.667 |
|  | B | 14 | 304.1 | -0.071 | 0.032 | 7.197 | 3.387 |
|  | C | 9 | 279.4 | -0.094 | 0.025 | 9.593 | 4.575 |
|  | D | 7 | 346.7 | -0.088 | 0.023 | 23.654 | 11.049 |
|  | E | 83 | 438.8 | -0.090 | 0.023 | 8.562 | 3.934 |
|  | A | 16 | 344.1 | -0.105 | 0.033 | 13.536 | 3.871 |
|  | D inshore | 14 | 431.8 | -0.109 | 0.031 | 19.598 | 4.840 |
|  | E | 10 | 336.6 | -0.136 | 0.027 | 14.266 | 4.090 |
|  | A | 8 | 433.5 | -0.125 | 0.026 | 24.564 | 5.819 |
|  | B | 96 | 550.2 | -0.121 | 0.025 | 10.940 | 2.848 |
|  | D | 16 | 236.5 | -0.205 | 0.077 | 6.975 | 2.521 |
|  |  | 14 | 234.9 | -0.196 | 0.080 | 9.439 | 3.233 |
|  |  | 10 | 278.8 | -0.242 | 0.087 | 43.657 | 13.670 |
|  |  | 8 | 281.7 | -0.179 | 0.074 | 46.948 | 15.414 |
|  |  | 293.3 | -0.221 | 0.072 | 11.450 | 4.110 |  |

1. A - Year, Depth, Depth ${ }^{2}$, Temperature, Temperature ${ }^{2}$, Latitude, Seasonal Cycle, Seasonal Cycle * Latitide; B - Year, Depth, Depth ${ }^{2}$, Latitude, Seasonal Cycle, Seasonal Cycle * Latitide; C - Year, Depth, Depth ${ }^{2}$, Temperature, Temperature ${ }^{2}$, Season; D - Year, Depth, Depth ${ }^{2}$, Season; E - Year, Season, Stratum, Season*Stratum


Figure X .

